Stochastic collocation for device-level variability analysis in integrated photonics

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We demonstrate the use of stochastic collocation to assess the performance of photonic devices under the effect of uncertainty. This approach combines high accuracy and efficiency in analyzing device variability with the ease of implementation of sampling-based methods. Its flexibility makes it suitable to be applied to a large range of photonic devices. We compare the stochastic collocation method with a Monte Carlo technique on a numerical analysis of the variability in silicon directional couplers. © 2016 Chinese Laser Press

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1. INTRODUCTION

Integrated photonics, and in particular, silicon photonics, is rapidly enabling complex photonic functions on a chip [1]. However, variability due to fabrication processes and operational conditions limits the complexity of the circuits that can be implemented [2,3]. A proper analysis of the effects of the variability in geometrical, electrical, and optical parameters on the performance of the photonic building blocks and circuits has become crucial. Indeed, the variability introduced by the variations of the manufacturing process is a primary source of degradation of larger photonic circuits, especially when wavelength-selective filters are implemented [4]. The primary functional parameters that are affected are the waveguide propagation constants (the effective index) and the coupling coefficients in coupling structures. The Monte Carlo (MC) method [5] is considered the standard approach for variability analysis, thanks to its accuracy and ease of implementation. Unfortunately, the MC analysis has a slow convergence rate, and it requires a large number of data points (simulations or measurements). Therefore, MC has a high computational cost, considering that accurate simulations of photonic devices can be time and resource intensive.

The generalized polynomial chaos (gPC) expansion has been applied in several domains as an efficient alternative to the classic MC method [6–8], and, recently, it has been proposed for the variability analysis of photonic devices [9,10].

The gPC-based modeling approach aims at expressing a stochastic process as a series of orthogonal basis functions with suitable coefficients and gives an analytical representation of the variability of the system on the random variables under consideration [11].

In this paper, we propose a stochastic collocation (SC) method as an efficient alternative to characterize photonic devices under the effect of uncertainty. The fundamental principle of the SC approach is to approximate the unknown stochastic solution by interpolation functions in the stochastic space. The interpolation is constructed by repeatedly solving (sampling) the deterministic problem at a predetermined set of nodes in the stochastic space. This approach offers similar high accuracy and efficiency as the stochastic gPC method, but, at the same time, it is easy to implement, such as sampling-based methods (e.g., MC approach).

We apply the SC method to analyze the variability of a key building block for silicon photonic circuits: the directional coupler (DC). This device is essential in the construction of wavelength filters, as it implements an arbitrary fractional 2×2 power coupling, but, at the same time, it is extremely sensitive to fabrication variations: a small shift in linewidth or thickness of the core can dramatically change the coupling coefficients.

The paper is organized as follows: Section 2 gives a general introduction of SC methods. It provides the essential mathematics knowledge that readers need to know to understand its application in the photonic domain. Section 3 uses a DC as an example to test the performance of SC in performing variability analysis of photonic devices. Section 4 draws the conclusions.

2. STOCHASTIC COLLOCATION METHODS

SC methods are based on interpolation schemes to compute stochastic quantities. The interpolation is constructed by repeatedly solving (sampling) the deterministic problem at a predetermined set of nodes in the stochastic space [12] (also defined as collocation points). Indeed, a stochastic process $\mathbf{Y}(\xi)$ can be expressed as

$$\mathbf{Y}(\boldsymbol{\xi}) = \sum_{i=1}^{Q} \mathbf{Y}(\boldsymbol{\xi}_i)L_i(\boldsymbol{\xi}), \tag{1}$$

where $\boldsymbol{\xi}$ denotes the N stochastic parameters and $\{L_i(\boldsymbol{\xi})\}_{n=1}^{N}$ represents the interpolation basis functions.

For a photonic device, the process $\mathbf{Y}$ could correspond to the functional parameters such as the waveguide propagating constants and the coupling coefficients in coupling devices. The stochastic variables ξ_i correspond to device properties affected in a stochastic way by fabrication and operational conditions (e.g., waveguide line, width, or temperature).

In Eq. (1), different types of interpolation schemes can be adopted (e.g., piecewise linear [12,13], Lagrange [11,14], or interpolation methods that belong to the general class of positive interpolation operators such as multivariate simplicial methods [15]). However, the key issue for this approach is the selection of the support nodes, such that when using the minimal number of nodes one achieves good approximation.

For example, if the Lagrange interpolation scheme is chosen, the element L_i in Eq. (1) for a 1D interpolation can be expressed as

$$L_i(\xi) = \prod_{i=1, i\neq j}^{Q} \frac{\xi - \xi_i}{\xi_j - \xi_i}, \tag{2}$$

where L_i is equal to 1 for $\xi = \xi_j$ and is equal to 0 for $\xi = \xi_i$. Next, for interpolation in multiple dimensions, a tensor-product approach can be used, and Eq. (1) becomes

$$\mathbf{Y}(\boldsymbol{\xi}) = \sum_{i_1=1}^{Q_{k_1}} \cdots \sum_{i_N=1}^{Q_{k_N}} \mathbf{Y}(\xi_{i_1}^{k_1}, \ldots, \xi_{i_N}^{k_N})(L_{i_1}^{k_1} \otimes \cdots \otimes L_{i_N}^{k_N}), \tag{3}$$

where ξ_i^k is the i-th node in the k-th direction, and the total number of nodes used in Eq. (3) is

$$Q = \prod_{n=1}^{N} Q_{k_n}. \tag{4}$$

As can be seen in Eq. (4), the number of nodes required by the full tensor product increases rapidly with the number of random parameters N. For example, if three random variables are considered and 10 collocation points are used for each parameter, a total of 1000 nodes are required by the full tensor product approach. Hence, the performance of the photonic device under study must be evaluated for 1000 different combinations of the random variables considered, leading to an expensive computational time.

The required number of nodes can be significantly reduced by adopting sparse grids in the stochastic space, based on the Smolyak algorithm [12,16–19]. By choosing the collocation points correctly, the Smolyak algorithm drastically reduces the total number of nodes used in the interpolation with respect to the full tensor product approach while preserving a high level of accuracy.

It is important to remark that the SC models are expanded using interpolation functions of independent random variables $\boldsymbol{\xi}$ [11]. In the general case of correlated random variables, decorrelation can be obtained via a variable transformation, such as the Nataf transformation [20] or the Karhunen–Loéve expansion [21].

The stochastic moments (mean, variance, …) can be computed utilizing analytical formulas and then efficiently, once the analytical form of the interpolation functions $\{L_i(\boldsymbol{\xi})\}_{n=1}^{N}$ has been decided. For example, if the random variables $\boldsymbol{\xi}$ are defined in the sample space Ω, the mean of $\mathbf{Y}(\boldsymbol{\xi})$ is defined as

$$\mu(\mathbf{Y}(\boldsymbol{\xi})) = \int_{\Omega} \mathbf{Y}(\boldsymbol{\xi})W(\boldsymbol{\xi})\mathrm{d}\boldsymbol{\xi}, \tag{5}$$

where $W(\boldsymbol{\xi})$ is the joint probability density function (PDF) of the random variables $\boldsymbol{\xi}$. Using Eq. (1) in Eq. (5) leads to

$$\mu(\mathbf{Y}(\boldsymbol{\xi})) = \int_{\Omega} \sum_{i=1}^{Q} \mathbf{Y}(\boldsymbol{\xi}_i)L_i(\boldsymbol{\xi})W(\boldsymbol{\xi})\mathrm{d}\boldsymbol{\xi}, \tag{6}$$

which depends only on the interpolation functions $L_i(\boldsymbol{\xi})$ and joint PDF $W(\boldsymbol{\xi})$. Note that, if the choice of the interpolation functions and probability measure does not allow an analytical computation of the stochastic moments like Eq. (6), an efficient numerical solution can be used (e.g., by MC analysis of the interpolation model or numerical integration). Finally, it is important to remark that it is not possible to define *a priori* the speed-up of a generic SC modeling technique compared with the MC method. Indeed, the number of nodes needed to compute an accurate SC model (which is directly related to the efficiency of SC methods, as described above) cannot be decided upfront because it depends on the following factors:

- the impact of the chosen random variables $\boldsymbol{\xi}$ on the variations of the stochastic process considered Y (dynamic stochastic processes require a higher number of collocation points);
- the interpolation scheme L_i adopted (the more powerful the interpolation scheme, the fewer nodes are needed);
- the sampling strategy adopted (efficient sampling strategy limit the number of collocation points used);
- the number of random variables considered (the higher the number of variables the more collocation points are needed).

However, it has been proven in the literature that, for a limited number of random variables (indicatively less than 10) SC methods are much more efficient with respect to the MC analysis (see [11,12,14]). For stochastic processes depending on a high number of random variables, the efficiency of SC methods is significantly reduced.

Two approaches can be used to increase the efficiency of an SC modeling technique. Using nested sampling schemes allows us to adaptively choose the collocation points (additional details are provided in Section 3.E and Appendix 5. B). Adopting adaptive sparse grids [16] reduces the nodes requirement, which is especially useful when a high number of random variables is considered. For a more detailed reference on SC methods, we refer the reader to [11,12,14,16].

3. DIRECTIONAL COUPLER EXAMPLE

A. Benchmark Description

We demonstrate the use of SC for integrated photonics through the analysis of a DC in a silicon photonics platform. Silicon photonics is rapidly gaining adoption due to its poten-

tial for large-scale integration and volume manufacturing. [22]. However, the same high-index contrast that enables dense integration also makes silicon photonic waveguides extremely sensitive to small imperfections in their geometry. Also, the high thermo-optic coefficient of silicon makes silicon photonic devices temperature sensitive. For instance, a change in linewidth and thickness of waveguide would noticeably vary the effective index (n_{eff}) of optical modes, resulting in a shift of 1 nm in the response of a wavelength-selective filter. This variation leads to performance degradation in devices such as DCs, Mach–Zehnder interferometers (MZIs) and ring resonators, and limits the number of devices that a single circuit can have.

Power coupling devices are essential parts for splitting and combining power in photonic circuits. Among them, DCs are widely used for their simplicity in layout and easy-to-understand operation. An advantage of DCs compared with other 2×2 couplers is that the coupling ratio can be continuously adjusted by choosing the length of the coupling section. Furthermore, DCs constitute the building blocks of many larger photonic devices such as rings, MZIs, and so on.

A DC consists of two parallel waveguides and connecting bend sections: the light in a single waveguide is mostly confined in the silicon core, but an exponentially decaying field extends into the cladding. When two waveguides are brought in proximity, the modes of the two waveguides couple and form two supermodes with opposite symmetry (an even and an odd mode). The beating of these supermodes translates in a sinusoidal power transfer from one waveguide to the next and back along the propagation axis z. The power coupling $K(z)$ in a DC can be expressed as [23]

$$K(z) = \sin^2(\kappa z + \kappa_0). \tag{7}$$

The power coupling consists of two parts: the contribution of the straight waveguide κz and of the bend section κ_0. The bend part will introduce an initial phase in the coupling term. Usually, the bend contribution is fairly small, and the power coupling is sinusoidally varying with the waveguide coupling section length z.

The rate of coupling is defined as the field coupling coefficient κ, which is determined by the geometry of the coupler cross section, such as the waveguide width, thickness, and gap between the waveguide cores.

Let us assume that, for simplicity, the two waveguides in a DC are identical. As a result, the straight section of the DC layout is defined by three parameters: the waveguide width w, thickness t, and gap g (Fig. 1). Furthermore, we assume that, in the lithography process, the centers of the two waveguides are located at the designed position. It is a good assumption for optical lithography techniques but might be less accurate for e-beam written devices. With this assumption, the sum of the gap g and 2× the half-waveguide width w is constant, as shown in Fig. 1. Therefore, in our example, we can describe the full geometry of the DC with only two parameters: w and t.

In this study, we will use the SC technique to find out how geometry variability influences the DC performance, namely, the coupling coefficient k. Indeed, due to the fabrication variations, the fabricated linewidth w, thickness t, and gap g are different on the value chosen during the design phase. To prove the robustness and modeling power of the proposed
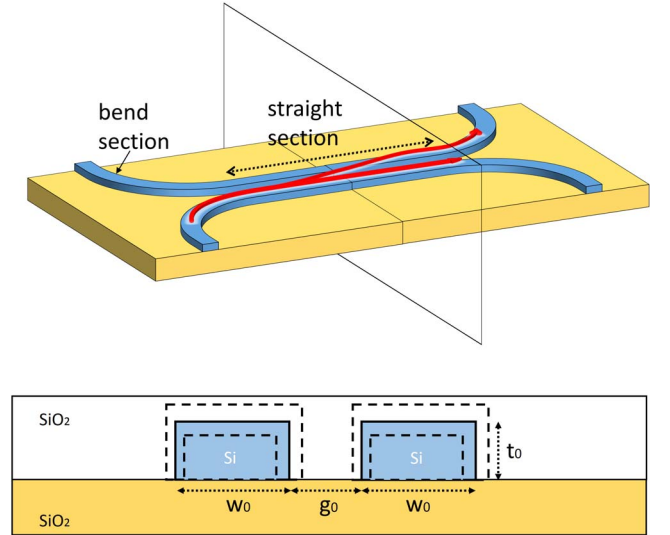

Fig. 1.   Upper plot shows the perspective view of a symmetric DC. Red arrows present the flow of light. Part of the light is coupled from bottom waveguide to the above one. Cross section is amplified in the lower plot. The mean width and thickness of the DC are w_0 and t_0, respectively. The width w and thickness t of the fabricated DC are indicated as dashed boxes. The refractive indexes are $n_{\text{si}} = 3.44$, $n_{\text{SiO}_2} = 1.45$.

approach, we assume the width w and thickness t of the DC as correlated random variables, rather than independent, following the Gaussian distribution. It is not an unrealistic assumption: thickness variations could induce over-etching on the sidewalls.

It is good to note that the SC methods can deal with random variables with arbitrary distribution. It is therefore not necessary that the t and w adhere to a Gaussian distribution.

B. Simulation Setup

According to the theory of supermodes, we can write the coupling coefficient κ as [23]

$$\kappa = \frac{\pi}{\lambda}(n_{\text{eff}\_o} - n_{\text{eff}\_e}), \tag{8}$$

where n_{eff_o} and n_{eff_e} are the effective index of asymmetrical and symmetrical supermodes in DC. For our silicon photonics devices, we assume the wavelength to $\lambda = 1.55\ \mu\text{m}$. Next, the nominal value of the width and thickness are $w_0 = 450$ nm and $t_0 = 220$ nm, respectively, while we fix the sum of width w and gap g at 650 nm.

To calculate κ of a given geometry, we define the DC structure accordingly and simulate n_{eff_o} and n_{eff_e} in the mode solver Fimmwave using its film matching mode (FMM) solver. For later performance comparisons, all simulations are performed on a computer with an Intel Core i5 2500 quad-core CPU clocked at 3.3 GHz and 8 GB of memory.

C. Problem Definition

As mentioned in Section 3.A, we considered the coupling coefficient K of a DC as a stochastic process depending on two correlated random variables with Gaussian PDFs: the width w and thickness t. Hence, the joint PDF of the two random variables considered is defined as

$$W_\eta = \frac{1}{2\pi \det(\mathbf{C})^{1/2}} \exp\left(-\frac{1}{2}(\eta - \mu)^T \mathbf{C}^{-1}(\eta - \mu)\right), \qquad (9)$$

where $\eta = [wt]^T$ is the vector of the correlated random variables considered, the vector $\mu = [w_0 t_0]^T$ contains the corresponding nominal values (mean values) w_0 and t_0, and the matrix $\mathbf{C}$ is the covariance matrix. The symbol $\det(\cdot)$ represents the matrix determinant operator. The covariance matrix is defined as

$$\mathbf{C} = \begin{bmatrix} (w_0\sigma_w)^2 & \rho w_0\sigma_w t_0\sigma_t \\ \rho w_0\sigma_w t_0\sigma_t & (t_0\sigma_t)^2 \end{bmatrix},$$

where the symbols σ_w and σ_t are the normalized standard deviations of the w and t, while ρ is the correlation coefficient of these two random variables. The correlation coefficient $|\rho| < 1$ denotes the strength of correlation: the random variables considered are independent if $\rho = 0$ and strongly correlated if $|\rho| = 1$. Note that, by describing this example in terms of normalized standard deviations, we make further analysis independent of the actual nominal values of our two random variables.

To validate the robustness of the proposed method, σ_w and σ_t are chosen equal to 2% and the correlation coefficient $\rho = 0.9$, which is a challenging example to study because the coupling coefficient is quite dynamic with respect to the parameters considered (see Fig. 2). The proposed method is discussed in detail in the following and summarized in Fig. 3.

D. Variable Transformation

Now, SC methods in the form of Eq. (1) deal with independent random variables, as described in Section 2. Hence, to fit the problem into the SC framework, first of all, it is necessary to express the coupling coefficient in two independent Gaussian random variables, starting from the correlated random variables η, defined by Eq. (9). As mentioned in Section 2, such decorrelation can be obtained via a variable transformation. Thanks to the Karhunen–Loéve expansion [21], it is possible to express the vector of correlated Gaussian random variables η in the vector of uncorrelated Gaussian random variables with zero mean and unit variance $\xi = [\xi_1, \xi_2]^T$ as

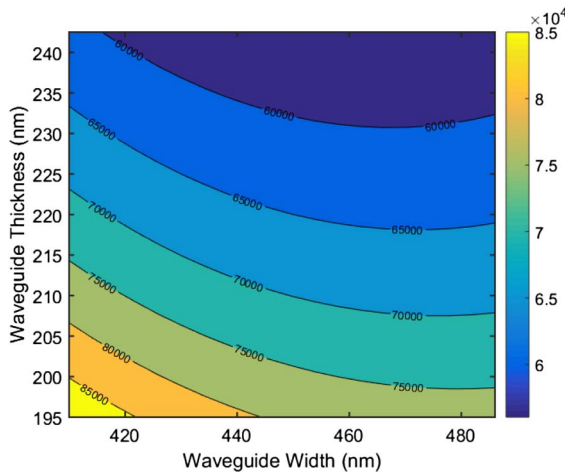$$\eta = \mu + \mathbf{V}\mathbf{E}^{1/2}\xi, \qquad (10)$$

where $\mathbf{E}$ and $\mathbf{V}$ are the diagonal matrix of the eigenvalues and the full matrix of the eigenvectors of the covariance matrix $\mathbf{C}$, respectively. Because uncorrelated Gaussian random variables are also independent, we have now expressed the coupling coefficient as a stochastic process, which depends on the pair of independent Gaussian random variables (ξ_1, ξ_2). An accurate description of the Karhunen–Loéve expansion for Gaussian random variables is given in Appendix 5.A.

E. SC Model Computation

To compute an SC in the form of Eq. (1), the first step is choosing the interpolation scheme: the Lagrange interpolation scheme is adopted in this example for its modeling power and ease of implementation. Next, a rule that guarantees a good quality of the approximation must be used to choose the collocation points for each random variable: ξ_1 and ξ_2. In this example, the Clenshaw–Curtis rule is adopted [17]: the collocation points for each random variable are the extrema of the Chebyshev polynomials. Now, the total number of nodes could be obtained using the full tensor products of the nodes chosen for each random variable, but it would not be efficient, as discussed in Section 2. Instead, the nodes are chosen over a sparse grid based on the Smolyak algorithm. Indeed, the adoption of the Smolyak algorithm allows building
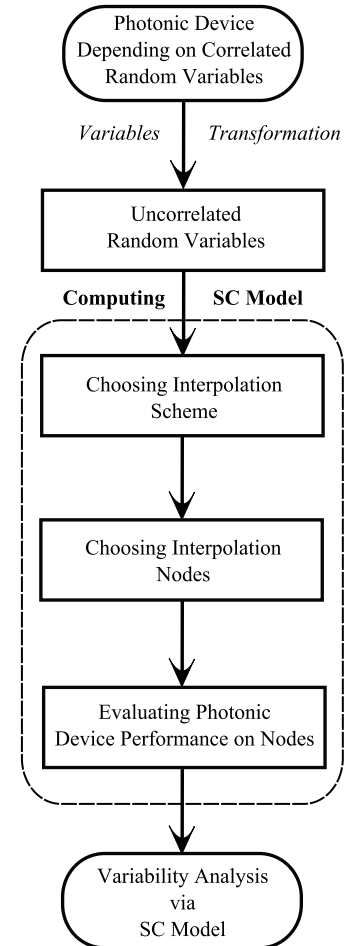

Fig. 2.   2D contour plot of field coupling coefficient versus waveguide width and thickness.

Fig. 3.   Flow chart of the proposed technique.

Fig. 5. Sampling points used to perform the MC analysis through direct Fimmwave simulations for the correlated random variables (w, t). The corresponding values for the independent random variables (ξ_1, ξ_2) are used to evaluate the SC model computed.
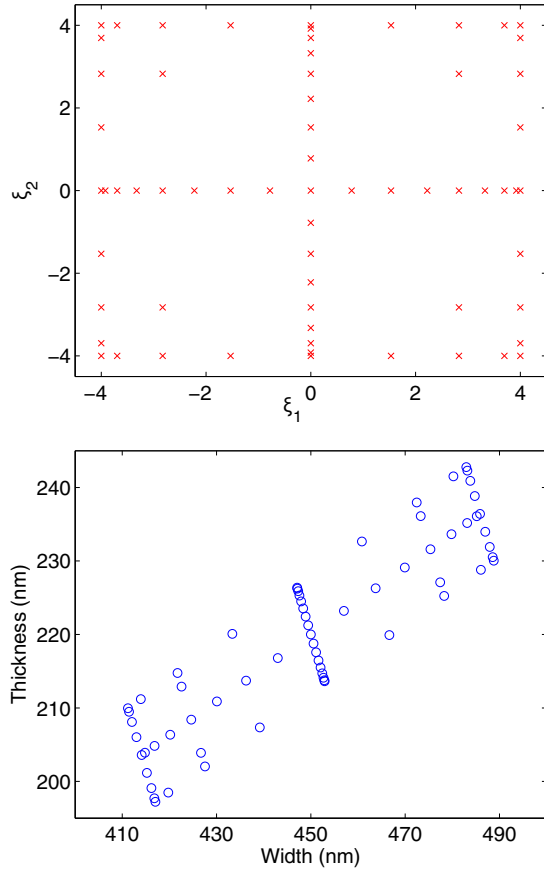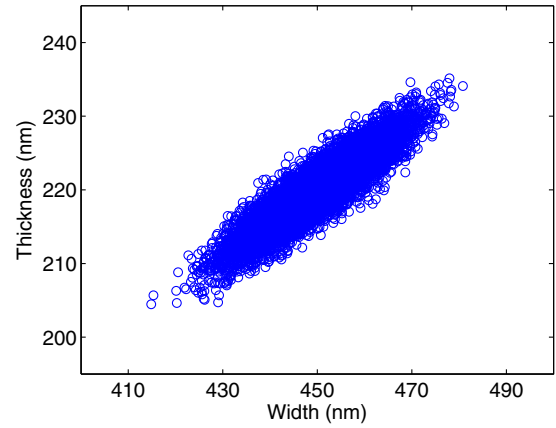

Fig. 4. Top: the red exes (×) represent the interpolation nodes for the normalized independent random variables ξ_1 and ξ_2 used to build the SC model. Bottom: the blue circles (°) are the corresponding values for the correlated random variables w and t used to compute the coupling coefficients in Fimmwave.

The proposed method shows an excellent accuracy compared with the classical MC analysis, as shown in Table 1, Figs. 6 and 7. In particular, the mean and the standard deviation of the coupling coefficient obtained employing the two methods are reported in Table 1: the relative error in the estimation of the mean and the standard deviation is only 9.0×10^{-5} and 5.6×10^{-3}, respectively. Apart from stochastic moments, more complicated functions of the stochastic process under study can be estimated: the probability density and cumulative distribution function (CDF) of κ obtained utilizing the two methods considered are in excellent agreement, as shown in Fig. 7.

From the field coupling coefficient, we can also easily derive performance parameters of a DC such as 3 dB coupling length from

Table 1. Performance Summary of SC and MC Simulation

	MC	SC
Mean value	65,160	65,166
S.t.d value	2616.9	2631.4

our SC model by using only a subset of all the collocation points given by the full tensor product [17]. Furthermore, the collocation points chosen by the Smolyak algorithm based on the Clenshaw–Curtis rule are nested: if additional nodes are required to accurately model the DC, the nodes already computed are kept in the new sparse grid, thus reducing the number of evaluation of the DC coupling coefficient. (See Appendix 5.B for additional details on the Smolyak algorithm.) As a result, only 65 collocation points (Fig. 4) are required to build the desired SC model, and the values of the coupling coefficient at the interpolation nodes are computed using the FMM solver Fimmwave.

F. Directional Coupler Variability Analysis
Finally, the variability analysis for the coupling coefficient of the DC under study is performed using an SC model depending on the pair of independent Gaussian random variables (ξ_1, ξ_2) and the results obtained are validated through comparison with an MC analysis based on the Fimmwave FMM solver on the DC cross section for the couple of correlated random variables (w, t). To compare the performance of the two methods, the same set of 10000 samples for the pair of correlated random variables (w, t) (see Fig. 5). The corresponding values for the independent random variables (ξ_1, ξ_2) are used to estimate the device variability features.
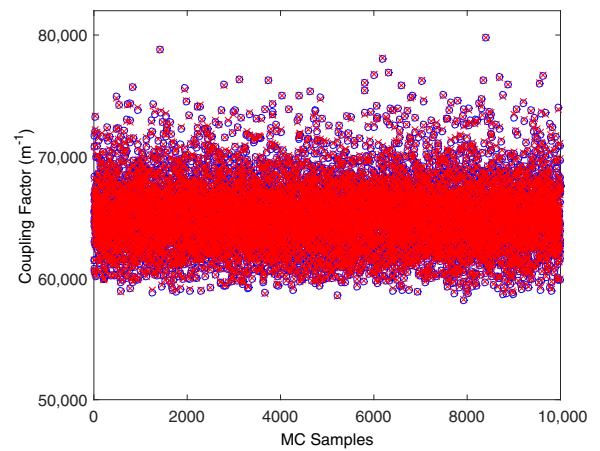

Fig. 6. Blue circles (°): coupling coefficient computed via the MC analysis for the 10000 (w, t) samples shown in Fig. 5. Red (×)-markers: corresponding values obtained by evaluating the SC model.
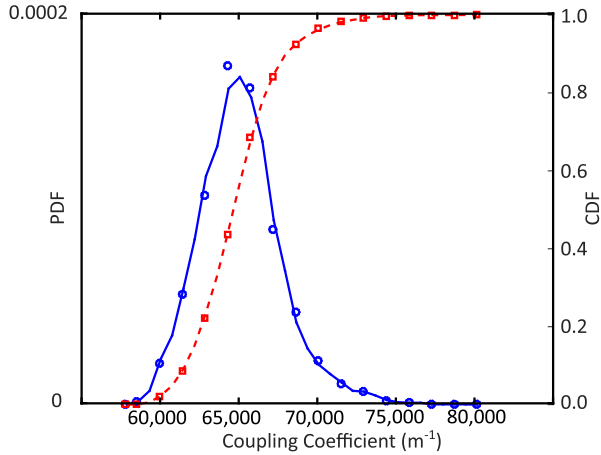
Fig. 7.   PDF and CDF of the coupling coefficient for $\lambda = 1.55$ µm. The blue solid and red dashed line are PDF and CDF obtained by means of the SC model, respectively, while the blue circles and red squares represent the same quantities computed by means of the MC analysis.

Fig. 8.   PDF and CDF of the 3 dB-coupling length for $\lambda = 1.55$ µm. The blue solid and red dashed line are PDF and CDF obtained by means of the SC model, respectively, while the blue circles and red squares represent the same quantities computed by means of the MC analysis.
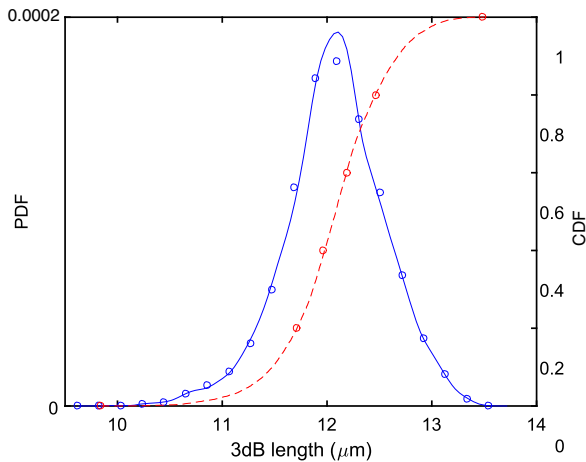
$$l_{3\,\mathrm{dB}} = \arcsin(\mathrm{sqrt}(0.5))/\kappa. \qquad (11)$$

As shown in Fig. 8, the proposed method shows excellent modeling accuracy and a good estimation of the PDF and the CDF of the 3 dB coupling length $l_{3\,\mathrm{dB}}$.

Furthermore, as presented in Table 2, the SC method has dramatically saved computational cost. Note that the SC method took a two-step procedure to perform the same variability analysis. Initially, SC required 65 simulations to compute the coupling coefficient at the collocation points. Next, we used the SC model over 10000 samples of the independent random variables in the MC method. Hence, the total computational time of the SC method is 8 min and 59 s, which represents a speed-up of a factor 146× with respect to the MC analysis performed in Fimmwave for the couple of correlated random variables (w, t), which required 21 h 53 min 14 s.

4. CONCLUSION

This paper has presented the application of a novel technique for the efficient variability analysis of photonic devices, such as DCs. It is based on the use of SC methods. Thanks to the flexibility in the choice of the interpolation schemes and the efficiency of sparse grid sampling to choose the collocation nodes for multiple dimension, the proposed approach is flexible and can be applied to study a broad range of photonic devices. The accuracy and efficiency of the proposed technique have been verified using comparison with the classic MC analysis for a pertinent numerical example, achieving a simulation speed-up of 146×.

Appendix A

KARHUNEN–LOÉVE EXPANSION AND CORRELATED GAUSSIAN RANDOM VARIABLES

Let us assume that the correlation matrix $\mathbf{C}^{N \times N}$ for the random variables $\boldsymbol{\eta}$ under study is symmetric and positive-definite. Then, $\mathbf{C}$ and can be diagonalized as

$$\mathbf{C} = \mathbf{V}\mathbf{E}\mathbf{V}^T \qquad (A1)$$

Thanks to Eq. (A1), Eq. (9) becomes

$$W_{\boldsymbol{\eta}} = \frac{1}{2\pi \det(\mathbf{E})^{1/2}} \exp\left(-\frac{1}{2}(\boldsymbol{\eta}-\boldsymbol{\mu})^T \mathbf{V}\mathbf{E}^{-1}\mathbf{V}^T(\boldsymbol{\eta}-\boldsymbol{\mu})\right). \quad (A2)$$

Hence, the Karhunen–Loéve expansion is a simple change of variables for correlated Gaussian random variables following the nondegenerate multivariate normal distribution in Eq. (9). Furthermore, it is possible to express the joint PDF Eq. (A2) with respect to a vector of independent Gaussian random variable $\boldsymbol{x}$, with zero mean and variance equal to $[\mathbf{E}_{ii}]_{i=1}^{N}$, as

$$W_{\boldsymbol{x}} = \frac{1}{2\pi \det(\mathbf{E})^{1/2}} \exp\left(-\frac{1}{2}\boldsymbol{x}^T \mathbf{E}^{-1}\boldsymbol{x}\right), \qquad (A3)$$

$$\boldsymbol{x} = \mathbf{V}^T(\boldsymbol{\eta}-\boldsymbol{\mu}). \qquad (A4)$$

Finally, the vector $\boldsymbol{x}$ can be written as

$$\boldsymbol{x} = \mathbf{E}^{1/2}\boldsymbol{\xi}, \qquad (A5)$$

Table 2.   Computation Time of SC and MC Simulation

Variability Analysis Technique		Simulator	Number of Points	Computation Time
	MC	Fimmwave FMM Solver	10000	21 h 53 min 14 s
	Stochastic modeling	Fimmwave FMM Solver	65	8 min 32 s
SC	MC using stochastic model	SC stochastic model	10000	27 s
	Total time			8 min 59 s

where $\boldsymbol{\xi}$ is a vector of normalized Gaussian random variables with zero mean and unitary variance. Equation (10) can be obtained by combining Eqs. (A4) and (A5).

SMOLYAK ALGORITHM

Let us express a stochastic process $\mathbf{Y}$ depending on one random variable ξ by means of the Lagrange interpolation scheme as [16]

$$\mathbf{U}(\xi) = \sum_{i=1}^{Q} \mathbf{Y}(\xi_i) L_i(\xi), \tag{A6}$$

where L_i is given by Eq. (2). The Q nodes can be chosen from a node distribution, which guarantees a good quality of the approximation (i.e., the extrema of the Chebyshev polynomials). Extending Eq. (A6) to the case of multiple random variables can be performed via tensor product, as has been shown in Section 2, and Eq. (A6) becomes

$$\mathbf{Y}(\boldsymbol{\xi};) = \mathbf{U}^{k_1} \otimes \cdots \otimes \mathbf{U}^{k_N}$$
$$= \sum_{i_1=1}^{Q_{k_1}} \cdots \sum_{i_N=1}^{Q_{k_N}} \mathbf{Y}(\xi_{i_1}^{k_1}, \dots, \xi_{i_N}^{k_N})(L_{i_1}^{k_1} \otimes \cdots \otimes L_{i_N}^{k_N}), \tag{A7}$$

where $\mathbf{U}^{k_j}$ represents the interpolation scheme in the form Eq. (A6) with respect to the random variable ξ_j, and N is the number of random parameters considered. The total number of nodes required to compute Eq. (A7) is given by the product of the nodes used for each random parameter, as shown in Eq. (4). Clearly, the required number of nodes grows quickly with respect to the number of parameters considered. Indeed, if only two nodes are used for each random variable, the total number of points required for a full-tensor product interpolation is $Q = 2^N$.

The Smolyak algorithm allows us to build multidimensional interpolation functions based on a minimal number of nodes by expressing the desired interpolation as a linear combination of tensor products. In particular, the property of the 1D interpolation is conserved for higher dimensions. Indeed, the sparse interpolant $A_{q,N}$ given by the Smolyak algorithm is

$$A_{q,N}(\boldsymbol{\xi}) = \sum_{q-N+1 \leq |\mathbf{k}| \leq q} (-1)^{q-|\mathbf{k}|} \binom{N-1}{q-\mathbf{k}} (\mathbf{U}^{k_1} \otimes \cdots \otimes \mathbf{U}^{k_N}), \tag{A8}$$

where $q - N$ is the order of interpolation, $A_{N-1,N} = 0$ and $\mathbf{k} = (k_1, \dots, k_N)$ with $|\mathbf{k}| = k_1 + \cdots + k_N$. Hence, the interpolation function is built by adding a combination of 1D interpolant of order k_j with the constraint that the total sum $|\mathbf{k}|$ across all parameters is between $q - N + 1$ and q. Note that k_j can be considered as the interpolation level along the j-th direction.

Let us denote Θ as the set of points utilized in the 1D function interpolation. According to Eq. (A8), the stochastic process $\mathbf{Y}$ must be computed at the nodes of the sparse grid $\mathbf{H}_{q,N}$ given by

$$\mathbf{H}_{q,N} = \bigcup_{q-N+1 \leq |\mathbf{k}| \leq q} \Theta_1^{k_1} \times \cdots \times \Theta_N^{k_N}. \tag{A9}$$

It is important to note that, by choosing a suitable node distribution, such as Chebyshev or Gauss–Lobatto points, the sets of collocation points Θ^k are nested. Hence, the sparse grid of order q contains all the nodes computed for the sparse grid of order $q - 1$, and the stochastic process $\mathbf{Y}$ must be evaluated only on few new collocation points.

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